

Root Mean Square Error of the Maximum Likelihood Estimate of the Parameters of Pareto Distribution

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ABSTRACT

Pareto distribution is one of the most important distributions in reliability work. In this paper, estimating the parameters of Pareto Distribution is presented using maximum likelihood estimation for different sample sizes. Computer simulation has been carried out to obtain results. Root mean square errors were computed for several values of the parameters. As the sample size increases, the root means square error of the maximum likelihood decreases.

Keywords: pareto distribution; root mean square error; maximum likelihood estimate; parameters; simulation

INTRODUCTION

The Pareto distribution is given by,

$$f(x) = \frac{\alpha k^{\alpha}}{x^{\alpha+1}}$$

where $x \ge k, k$ is the lower bound is a scale parameter of the data and α is the shape parameter. This is sometimes called a power law distribution. It has a property that large numbers are rare, but smaller numbers are more common, and thus it is more common for a person to make a small amount of money versus a large amount of money [1].

This distribution is a good model for personal incomes, the assets of firms, the sizes of cities and the number of firms in various industries. Pareto distribution is a powerful tool for modeling a variety of real-life phenomena in social sciences, and actuarial. It is named after the Italian economist Vilfredo Pareto (1848-1923), who developed the distribution in the 1890's as a way to describe the allocation of wealth in society. He famously observed that 80% of society's wealth was controlled by 20% of its population, a concept now known as the "Pareto Principle" or the "80-20 Rule" (2).

The maximum likelihood estimates of the two parameters

are
$$\hat{k} = \min\{x_i\}$$
 and $\hat{\alpha} = \frac{n}{\sum_{i=1}^n \log\left(\frac{x_i}{k_j}\right)}$

The Root mean square error (RMSE) of an estimator of a population parameter is the square root of the mean square error (MSE). The mean square error is defined as the expected value of the square of the difference between the estimator and the parameter.

In this paper, we used computer simulation to compute the root mean square error of the maximum likelihood estimate of the parameters of Pareto Distribution.

THE ALGORITHM

To compute the root, mean square error, the following are the steps.

- (1) Generate n random numbers ξ_i from 0 to 1.
- (2) Generate n random variates $x_i = \frac{k}{(1-\xi_i)^{\frac{1}{\alpha}}}$.

(3) Compute
$$\widehat{k_i} = min \{x_i\}$$
 for j=1 to 100.

(4) Compute
$$\widehat{\alpha}_{j} = \frac{n}{\sum_{i=1}^{n} log\left(\frac{x_{i}}{\overline{k_{j}}}\right)}$$
 for j=1 to 100.

- (5) Compute the means of the estimators.
- (6) Compute the variances of the estimators.
- (7) Compute the root mean square error using the following formulas:

a.
$$RMSE(k) = \sqrt{\frac{1}{100} \sum_{j=1}^{100} (\hat{k}_j - k)^2}$$

b. $RMSE(\alpha) = \sqrt{\frac{1}{100} \sum_{j=1}^{100} (\hat{\alpha}_j - \alpha)^2}$

THE RESULTS

As the number of sample size n increases, we can observe the trends of the means of the Maximum Likelihood Estimates of the parameters k = 5, ($\alpha = 2$, $\alpha = 5$, and $\alpha =$ 10 approach to the true values (Tables 1, 2 and 3). This supports that maximum likelihood estimates of the parameters are unbiased estimators, that is $\hat{\theta}$ is an unbiased estimator of θ if $E(\hat{\theta}) = \theta$ [3]. These are pictured by the graphs as seen in Figures 1, 2 and 3.

| TABLE 1: Means of the Parameters (k=5, α |
|--|
|--|

| Number of samples (n) | <i>k</i> = 5 | $\alpha = 2$ |
|-----------------------|--------------|--------------|
| 5 | 5.55 | 3.347 |
| 10 | 5.266 | 2.597 |
| 30 | 5.087 | 2.146 |
| 50 | 5.051 | 2.061 |
| 100 | 5.022 | 2.034 |
| 500 | 5.077 | 2.013 |
| 1000 | 5.003 | 1.98423 |

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FIGURE 1: Graph of the means of the parameters k=5, $\alpha=2$.

| TABLE 2. Means o | f the Parameters | k = 5 | $\alpha = 5$ |) |
|--------------------|-------------------|---------|--------------|----|
| TADLE 2. Micallo U | i the i arameters | (n - J) | , u — J | J٠ |

| Number of samples (n) | <i>k</i> = 5 | <i>a</i> = 5 |
|-----------------------|--------------|--------------|
| 5 | 5.1988 | 7.8101 |
| 10 | 5.1079 | 6.3232 |
| 30 | 5.0371 | 5.3566 |
| 50 | 5.0194 | 5.1649 |
| 100 | 5.01 | 5.1399 |
| 500 | 5.0019 | 5.043 |
| 1000 | 5.0011 | 4.8866 |



FIGURE 2: Graph of the means of the parameters k=5, $\alpha=5$.

TABLE 3: Means of the Parameters ($k = 5, \alpha = 10$).

| Number of samples (n) | <i>k</i> = 5 | <i>a</i> = 10 |
|-----------------------|--------------|---------------|
| 5 | 5.102 | 16.27 |
| 10 | 5.045 | 13.02 |
| 30 | 5.015 | 10.6 |
| 50 | 5.009 | 10.54 |
| 100 | 5.005 | 10.2 |
| 500 | 5.001 | 10.09 |
| 1000 | 5 | 10.01 |



FIGURE 3: Graph of the means of the parameters k=5, $\alpha=10$.

Similarly, the variances for k = 5, $(\alpha = 2, \alpha = 5, \alpha = 10)$ converge to zero as the sample sizes increase (Tables 3, 4 and 5). This can be appreciated by viewing the graphs as depicted in Figures 3, 4 and 5. With these results, we can say that the estimators are consistent. Estimators are consistent if $\lim_{n\to\infty} Var(\hat{\theta}) = 0$ [3].

TABLE 4: Variances of the Parameters ($k = 5, \alpha = 2$).

| Number of samples (n) | k = 5 | <i>a</i> = 2 |
|-----------------------|----------|--------------|
| 5 | 0.389 | 4.822 |
| 10 | 0.079 | 1.604 |
| 30 | 0.006 | 0.154 |
| 50 | 0.003 | 0.093 |
| 100 | 0.001 | 0.047 |
| 500 | 6.00E-05 | 0.012251 |
| 1000 | 9.00E-06 | 0.003887 |



FIGURE 4: Graph of the variances of the parameters k=5, $\alpha=2$.

TABLE 5: Variances of the Parameters ($k = 5, \alpha = 5$).

| Number of samples (n) | k = 5 | <i>a</i> = 5 |
|-----------------------|----------|--------------|
| 5 | 0.035269 | 21.40974 |
| 10 | 0.015394 | 4.318826 |
| 30 | 0.001485 | 1.271762 |
| 50 | 0.003 | 0.590717 |
| 100 | 9.03E-05 | 0.23629 |
| 500 | 3.35E-06 | 0.055196 |
| 1000 | 1.24E-06 | 0.021706 |





TABLE 6: Variances of the Parameters ($k = 5, \alpha = 10$).

| Number of samples (n) | k = 5 | <i>a</i> = 10 |
|-----------------------|----------|---------------|
| 5 | 0.00966 | 139.776 |
| 10 | 0.00206 | 21.0265 |
| 30 | 0.00024 | 3.89196 |
| 50 | 0.0001 | 2.07936 |
| 100 | 2.50E-05 | 0.9174 |
| 500 | 1.10E-06 | 0.33507 |
| 1000 | 1.30E-07 | 0.08682 |



FIGURE 6: Graph of the variances of the parameters k=5, $\alpha=10$.

Accordingly, the root means square errors of the maximum likelihood estimates k = 5, ($\alpha = 2$, $\alpha = 5$, $\alpha = 10$) converge to zero (Tables 7, 8 and 9). Graphs of these convergences are pictured in Figures 7, 8 and 9 respectively.

TABLE 7: Root Mean Square Error of the Parameters $(k = 5, \alpha = 2)$.

| Number of samples (n) | k = 5 | <i>a</i> = 2 |
|-----------------------|---------|--------------|
| 5 | 0.83 | 2.567 |
| 10 | 0.385 | 1.395 |
| 30 | 0.114 | 0.398 |
| 50 | 0.073 | 0.31 |
| 100 | 0.046 | 0.223 |
| 500 | 0.01 | 0.111 |
| 1000 | 0.00424 | 0.06401 |



FIGURE 7: Graph of the RMSE of the parameters k=5, $\alpha=2$.

TABLE 8: Root Mean Square Error of the Parameters $(k = 5, \alpha = 5)$.

| Number of samples (n) | k = 5 | <i>a</i> = 5 |
|-----------------------|--------|--------------|
| 5 | 0.2728 | 5.4059 |
| 10 | 0.164 | 2.492 |
| 30 | 0.0534 | 1.2367 |
| 50 | 0.026 | 0.8419 |
| 100 | 0.0138 | 0.5461 |
| 500 | 0.0026 | 0.429 |
| 1000 | 0.0016 | 0.3262 |



FIGURE 8: Graph of the RMSE of the parameters k=5, α=5.

TABLE 9: Root Mean Square Error of the Parameters $(k = 5, \alpha = 10)$.

| Number of samples (n) | k = 5 | <i>a</i> = 2 |
|-----------------------|----------|--------------|
| 5 | 0.141 | 13.33 |
| 10 | 0.064 | 5.417 |
| 30 | 0.022 | 2.054 |
| 50 | 0.014 | 1.532 |
| 100 | 0.007 | 0.973 |
| 500 | 0.001 | 0.825 |
| 1000 | 5.00E-04 | 0.576 |



FIGURE 9: Graph of the RMSE of the parameters k=5, $\alpha=10$.

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CONCLUDING REMARKS

Simulation results show that the root mean square errors of the maximum likelihood estimates of the parameters decrease to zero, as we increase the sample size. The Mean Absolute Error of the maximum likelihood estimate of the parameters is a potential study, and so with the other methods of estimating the parameters of the Pareto Distribution. It is suggested to conduct comparative evaluation. This is very interesting since the Root Mean Square Error and the Mean Absolute Error are defined differently, we should expect the results to be different [4].

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